

# Lecture 16

Note Title

2018-02-08

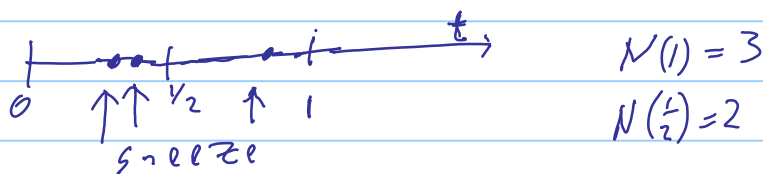
## Poisson processes

Goal: random model for the following as a function of time

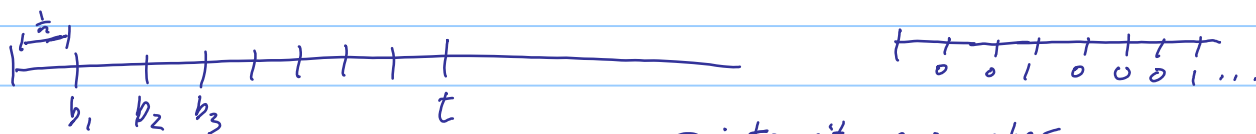
- Cars passing on highway
- earthquakes
- sneezes
- 
- 

$$N(t), t \geq 0 \quad (t \in \mathbb{R}_+)$$

$N(t)$  = number of occurrences in  $[0, t]$ .



First attempt: discrete model. Discretize  $\mathbb{R}_+$



Let  $b_i \stackrel{\text{indep}}{\sim} \text{Bernoulli}\left(\frac{\lambda}{n}\right)$  ← intensity parameter

$$\text{Let } N(t) = \sum_{i=1}^{nt} b_i$$

Q: ① What is dist of  $N(t)$ ?

② What is dist of waiting time to first event?  
 $T = n \cdot \min \{i : b_i = 1\}$

A:

$$\textcircled{1} N(t) \sim \text{Bin}\left(\frac{\lambda}{n}, nt\right) \quad (\text{assuming } t \in \mathbb{Z}_+)$$

$$\textcircled{2} T = \frac{1}{n} \text{Geom}\left(\frac{\lambda}{n}\right)$$

Q: What happens to these distributions as  $n \rightarrow \infty$ .

$$\begin{aligned} \textcircled{1} P(N(t) = k) &= P\left(\text{Bin}\left(\frac{\lambda}{n}, nt\right) = k\right) \quad \text{for } t \in \mathbb{Z}_+ \\ &= \binom{nt}{k} \cdot \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{nt-k} \\ &= \frac{(nt)!}{k!(nt-k)!} \cdot \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{nt-k} \\ &= \frac{1}{k!} \cdot \frac{nt \cdot (nt-1) \cdot (nt-2) \cdots (nt-k+1)}{n^k} \cdot \lambda^k \cdot \left(1 - \frac{\lambda}{n}\right)^{nt-k} \end{aligned}$$

$$\Rightarrow \lim_{n \rightarrow \infty} P(N(t) = k) = \frac{1}{k!} \cdot t^k \lambda^k e^{-\lambda t} = \frac{(\lambda t)^k}{k!} e^{-\lambda t}$$

Thus  $N(t) \xrightarrow{\text{dist}} \text{Poisson}(\lambda t)$

Def We say  $X \sim \text{Poiss}(\lambda)$  if

$$P(X = k) = \frac{\lambda^k}{k!} e^{-\lambda} \quad k = 0, 1, \dots$$

$$\begin{aligned} \textcircled{2} P(T > t) &= P\left(\frac{1}{n} \text{Geom}\left(\frac{\lambda}{n}\right) > t\right) \\ &= P\left(\text{Geom}\left(\frac{\lambda}{n}\right) \geq tn\right) \\ &= \sum_{k=tn}^{\infty} \left(1 - \frac{\lambda}{n}\right)^k \cdot \frac{\lambda}{n} \quad \text{assuming } t \in \mathbb{Z}_+ \\ &= \frac{\lambda}{n} \frac{\left(1 - \frac{\lambda}{n}\right)^{tn}}{1 - \left(1 - \frac{\lambda}{n}\right)} \\ &= \left(1 - \frac{\lambda}{n}\right)^{tn} \end{aligned}$$

$$\text{Thus } \lim_{n \rightarrow \infty} P(T > t) = \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^{tn} = e^{-\lambda t}$$

$\Rightarrow T \xrightarrow{\text{dist}} \text{Exp}(\lambda)$

### Continuous case

We follow this intuition to define properties of a Poisson process.

Let  $N(t)$ ,  $t \geq 0$ , be a Poisson process w/ intensity  $\lambda$ . Then

① For  $t > s$ ,  $N(t) - N(s) \sim \text{Poisson}(\lambda t) = \text{Pois}(\lambda t)$

② The waiting time  $T =$  time it takes  $\sim \text{Exp}(\lambda)$   
between events for  $N(t)$  to increase  
by 1

③ For  $t > s \geq u > v$   
 $N(t) - N(s)$  is indep of  $N(u) - N(v)$  (independent increments)